

Quantification of length-bias in screening trials with covariate-dependent test sensitivity

Dissertation Research

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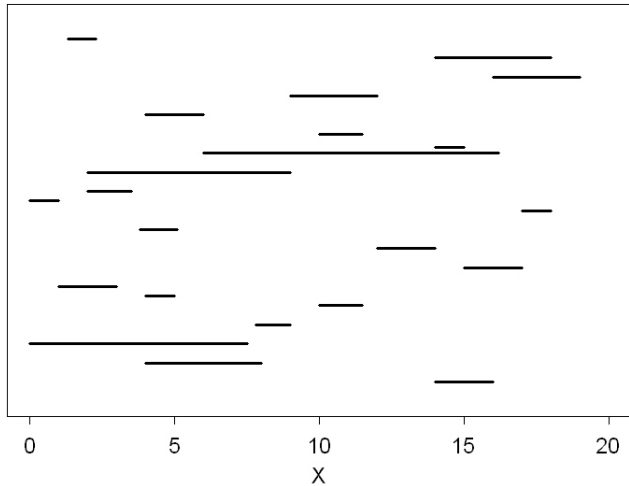
Outline

1. Length Bias/Inspection Paradox
2. Background: randomized screening trials
3. Length bias model in periodic screening
4. Results
5. Conclusion

1. Length Bias or Inspection Paradox

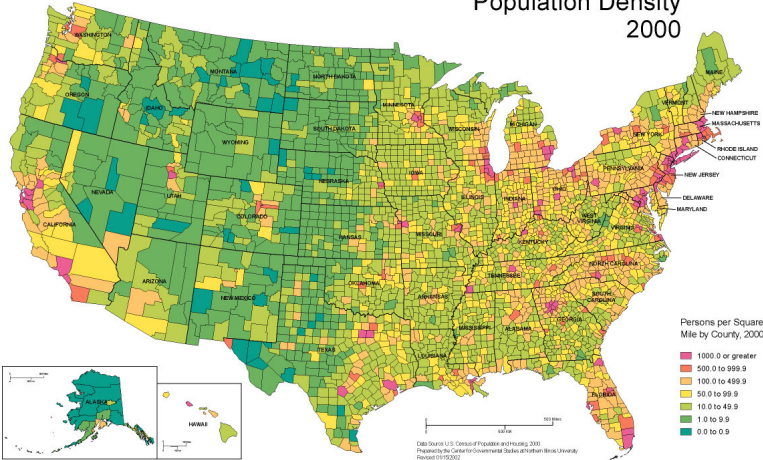
- Sampling bias
- Probability of being sampled is proportional to size
 - lengths of yarn
 - traffic
 - gas lines
 - warranty database
 - family size
 - residential sampling
 - disease screening
 - proteomics (mass spectrometry)

Length bias



Cluster size bias

Population Density 2000

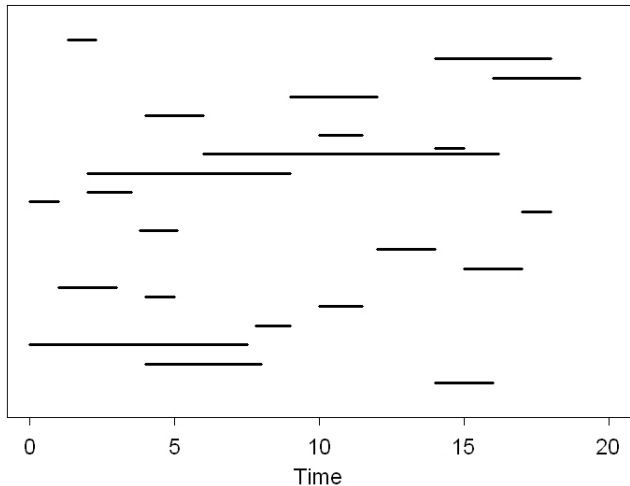


Length Bias or Inspection Paradox

- Industrial data
 - sampled variable is estimated
 - string
 - warranty
 - protein
 - true size/length is observable
- Disease/health data
 - sampled variable is correlated to estimable variable (Y, Z)
 - Hospital data: sample Stay, estimate \$
 - Disease data: sample Pre-clinical, estimate Clinical duration

Length Bias: Example

Sample hospital stays to estimate final bill (\$)



2. Randomized screening trials

- Screening mechanisms are used to detect disease early
- Participants are randomized to one of two arms
 - Study Arm -offered periodic screening.
 - Control Arm -standard medical care.
- Outcome of interest
 - disease specific mortality rates
 - extension of survival time

Randomized Screening Trials

- Disease progression model for unscreened and screen-detected cases if screening benefit exists

Unscreened Case:

| ← Preclinical Duration → | ← Clinical Duration → |

Screen-detected case at X:

X
|LeadTime | ← Clinical Duration → | Benefit|
| ← Preclinical Duration → |

Randomized Screening Trials



- Survival since Dx in *unscreened* arm is an unbiased estimate of clinical duration
- Average lead time can be estimated using time since entrance into study
 - b/w arm difference in average survival time since entrance into study: $\Delta_{t_0} = \overline{benefit}$
 - b/w arm difference in average survival time since Dx: $\Delta_{D_x} = \overline{lead} + \overline{benefit}$
 - $\overline{lead} = \Delta_{D_x} - \Delta_{t_0}$

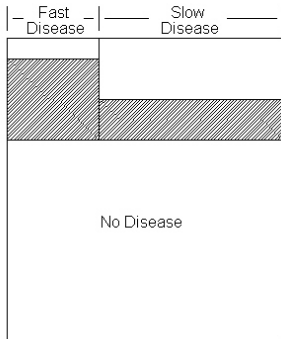
Randomized Screening Trials

- Survival time since entry into the *screened* arm confounds average benefit time, and average length bias
- Length bias:
 - when the probability of selection is proportional to the size (length) of the variable being sampled
 - Zelen (1976) recognized this phenomenon in screening. "Cases found by screening tend to be less advanced... slower-growing disease"
 - of particular concern when pre-clinical duration is positively correlated with clinical duration

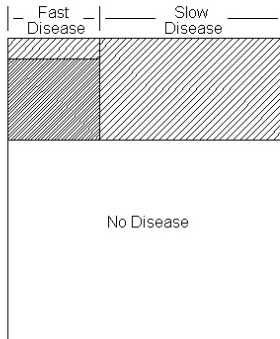
Distribution of disease and detection in the two study arms

Screening test sensitivity = 1.0.
Average sojourn time for 'fast disease'
is less than time between screens

Clinically Detected 
Screen Detected 



Control Arm



Screened Arm

3. Length bias in periodic screening

A. Single Sample

- (Y, Z) denote preclinical and clinical duration of disease in the diseased general population.
- $E(Y) = \mu_Y, E(Z) = \mu_Z, SD(Y) = \sigma_Y, SD(Z) = \sigma_Z, k.$
- (Y^*, Z^*) denote the same in the length-biased sample. Cox and Lewis (1972) show:

$$f_{Y^*}(y) = \frac{y}{\mu_Y} f_Y(y)$$

- Schotz & Zelen, 1971; Kafadar & Prorok, 2005

$$f_{Y^*, Z^*}(y, z) = \frac{y}{\mu_Y} f_{Y, Z}(y, z)$$

Length bias in periodic screening

B. Successive Sampling

- Goal: Estimate $E(Z^*)/E(Z)$:
 - Find density function of Y_k^* : $f_{Y_k^*}(y) = g_k(y)f_Y(y)$
 - Calculate $E(Y_k^*) = \int_0^\infty yg_k(y)f_Y(y)dy$.
 - Calculate expected sojourn time for screen-detected cases:

$$E(Y^*) = \sum_{k=1}^K \beta(1 - \beta)^{k-1} E(Y_k^*)$$

- Calculate $E(Z^*)/E(Z)$:

$$E(Z^*) = \sum_{k=1}^K \beta(1 - \beta)^{k-1} E(g_k(Y)Z)$$

Length bias in periodic screening

Estimate $E(Z^*)/E(Z)$ under a number of conditions:

- Distribution of pre-clinical duration
 - exponential, gamma, lognormal, weibull
- Correlation between pre-clinical and clinical duration, ρ
- Proportions of fast versus slow disease, ϕ
- moments of fast and slow disease ($\mu_\phi, \sigma_\phi^2, \mu_{(1-\phi)}, \sigma_{(1-\phi)}^2$)
- Ratio of time between screens, δ , and mean preclinical duration
- Variable test sensitivity: $\beta(y)$

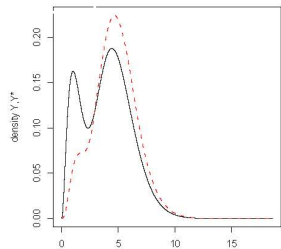
Disease mixture scenarios:

- **A:** Fast ($\mu_Y = 1.5, \mu_Z = 2$)
& Slow ($\mu_Y = 5, \mu_Z = 6$)
- **B:** Fast ($\mu_Y = 1, \mu_Z = 1.5$)
& Moderate ($\mu_Y = 3, \mu_Z = 2$)
- **C:** Moderate ($\mu_Y = 3, \mu_Z = 2$)
& Slow ($\mu_Y = 7, \mu_Z = 9$)
- **E:** Very Fast ($\mu_Y = 0.5, \mu_Z = 0.5$)
& Moderate ($\mu_Y = 3, \mu_Z = 2$)

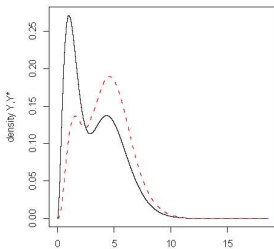
Scenario A : $\delta = 3$ $\beta = 0.7$

Fast/Slow

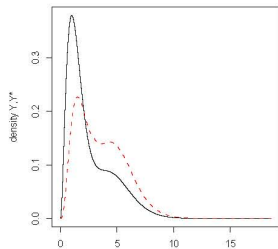
Y, Y^* : $\Phi = 0.3$



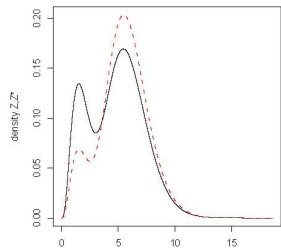
Y, Y^* : $\Phi = 0.5$



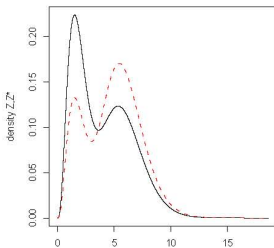
Y, Y^* : $\Phi = 0.7$



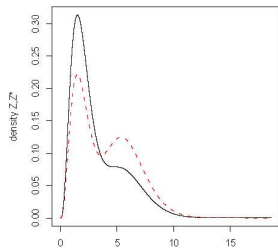
Z, Z^* : $\Phi = 0.3$

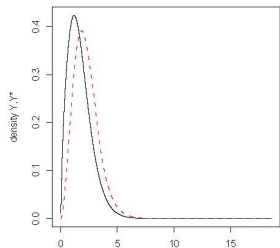
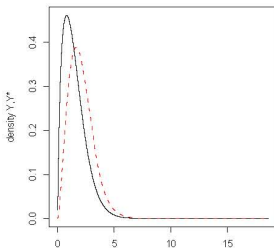
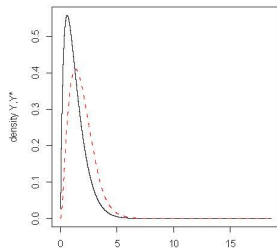
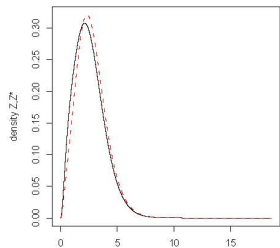
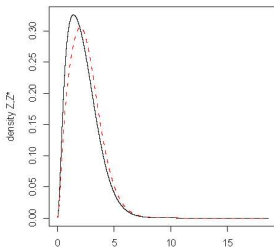
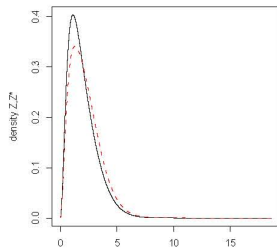


Z, Z^* : $\Phi = 0.5$



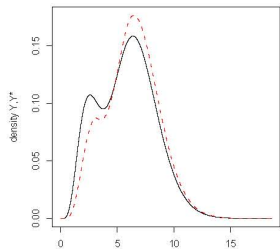
Z, Z^* : $\Phi = 0.7$



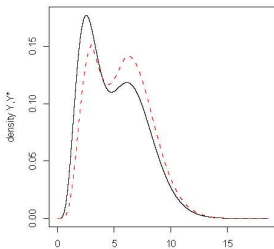
Scenario B : $\delta = 3$ $\beta = 0.7$ **Fast/Moderate** **$Y, Y^* : \Phi = 0.3$**  **$Y, Y^* : \Phi = 0.5$**  **$Y, Y^* : \Phi = 0.7$**  **$Z, Z^* : \Phi = 0.3$**  **$Z, Z^* : \Phi = 0.5$**  **$Z, Z^* : \Phi = 0.7$** 

Scenario C : $\delta = 3$ $\beta = 0.7$

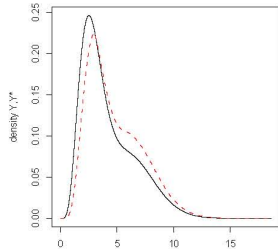
Moderate/Slow Y, Y^* : $\Phi = 0.3$



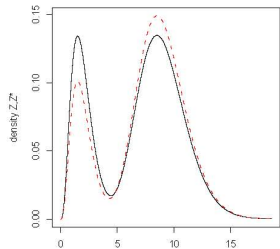
Y, Y^* : $\Phi = 0.5$



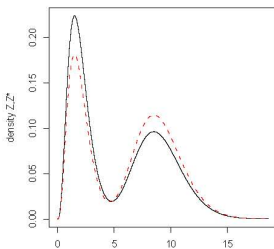
Y, Y^* : $\Phi = 0.7$



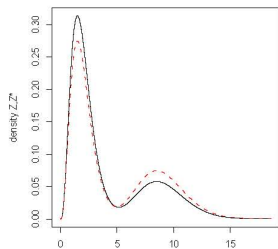
Z, Z^* : $\Phi = 0.3$



Z, Z^* : $\Phi = 0.5$



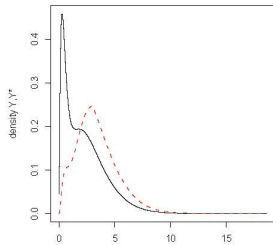
Z, Z^* : $\Phi = 0.7$



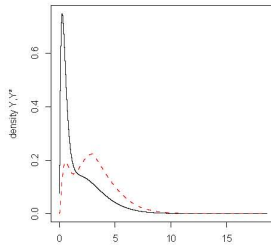
V. Fast/Moderate

Scenario E : $\delta = 3$ $\beta = 0.7$

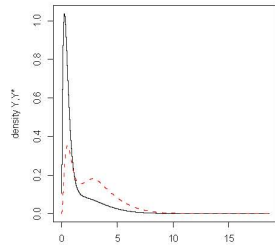
Y, Y^* : $\Phi = 0.3$



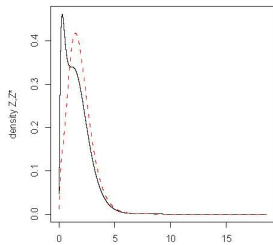
Y, Y^* : $\Phi = 0.5$



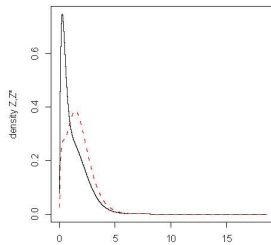
Y, Y^* : $\Phi = 0.7$



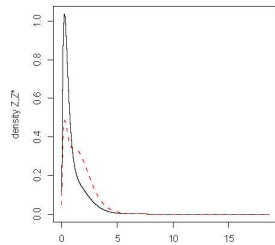
Z, Z^* : $\Phi = 0.3$



Z, Z^* : $\Phi = 0.5$

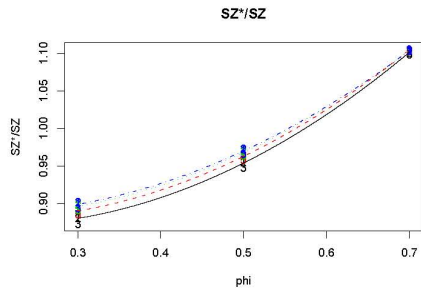
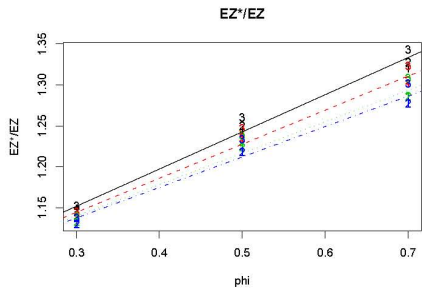
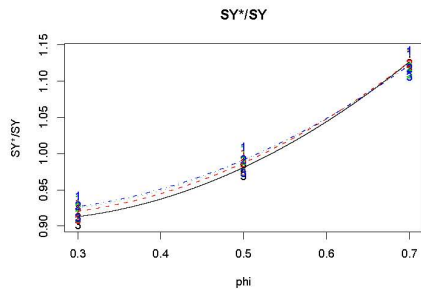
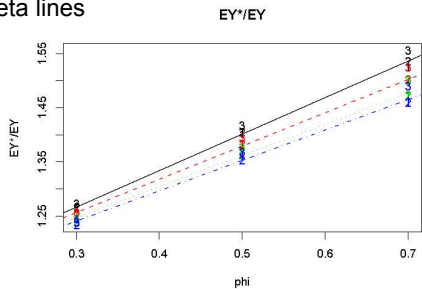


Z, Z^* : $\Phi = 0.7$



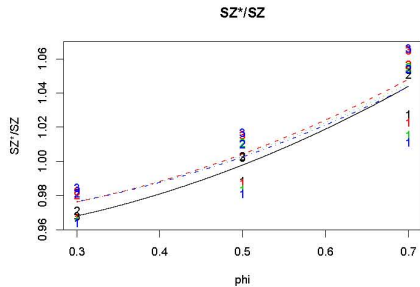
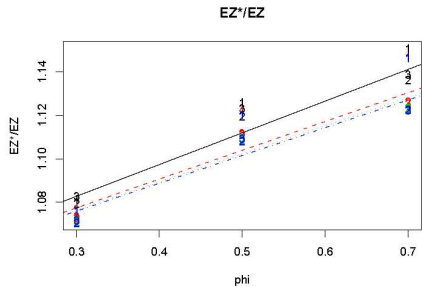
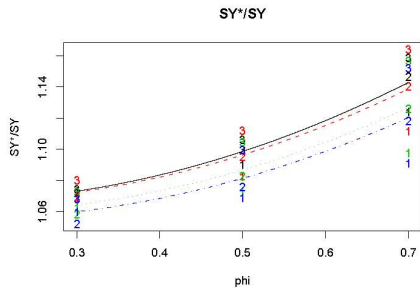
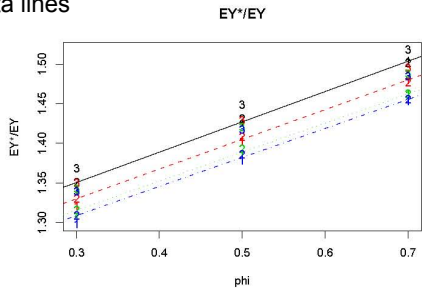
Fast/Slow Beta lines

Scenario A



Fast/Moderate Beta lines

Scenario B

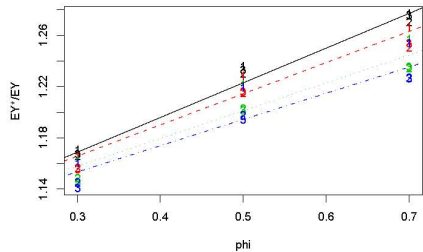


Moderate/Slow

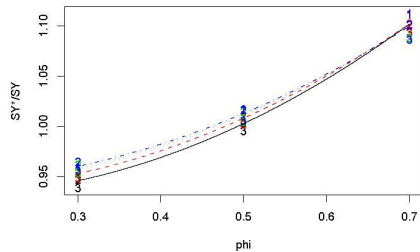
Scenario C

Beta lines

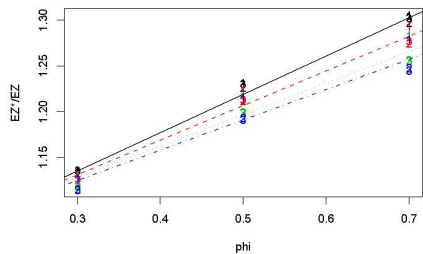
EY*/EY



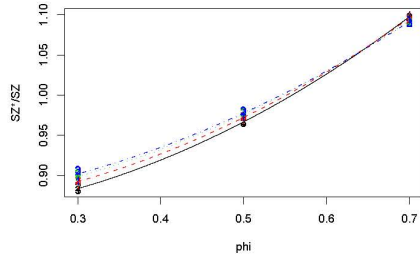
SY*/SY



EZ*/EZ



SZ*/SZ

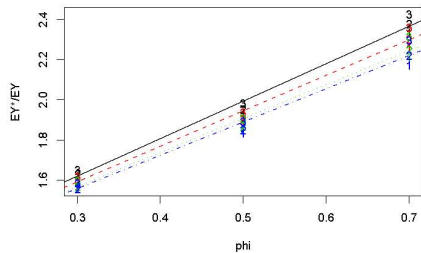


Scenario E

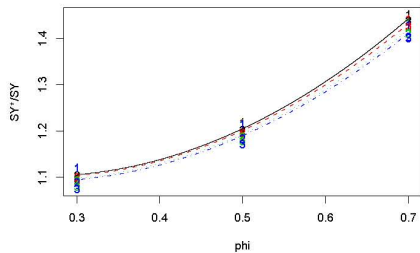
V.Fast/Moderate

Beta lines

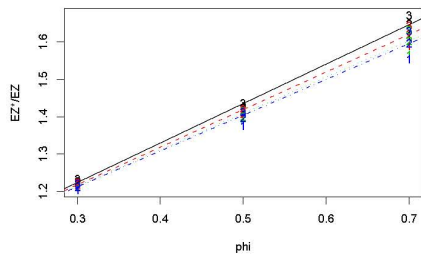
EY*/EY



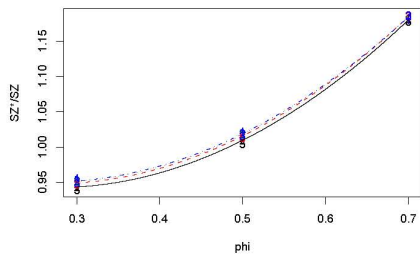
SY*/SY



EZ*/EZ

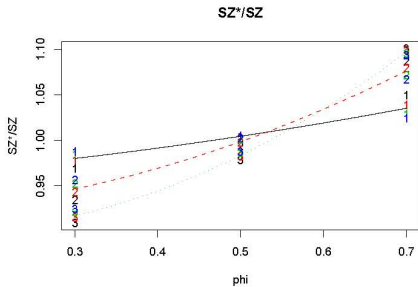
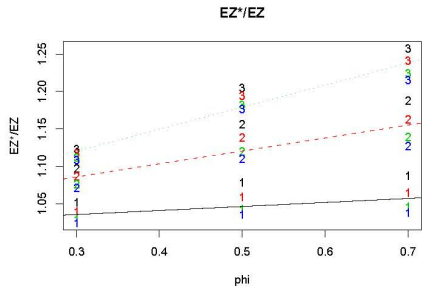
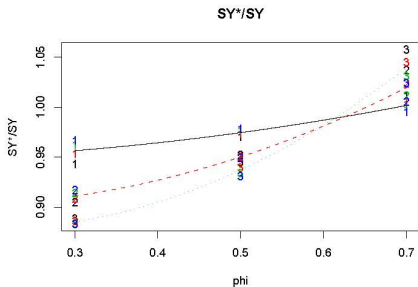
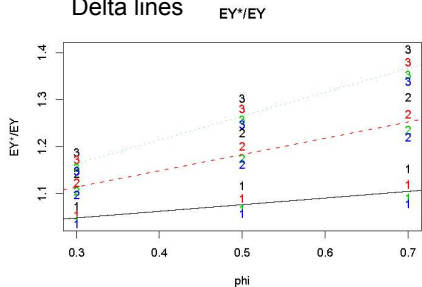


SZ*/SZ



Scenario A (5 periodic screens)

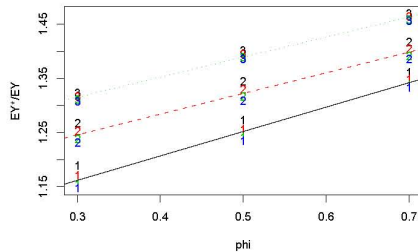
Fast/Slow
Delta lines



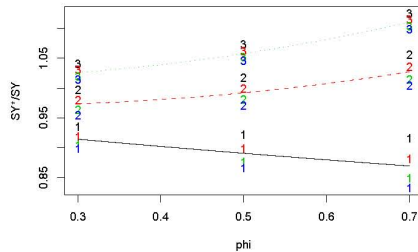
Fast/Moderate Delta lines

Scenario B (5 periodic screens)

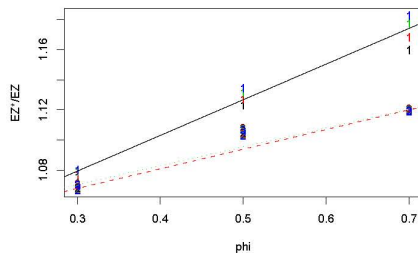
EY^*/EY



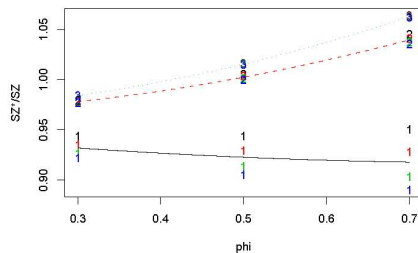
SY^*/SY



EZ^*/EZ



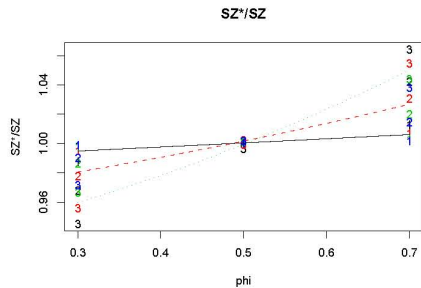
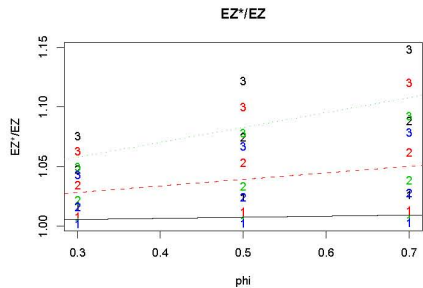
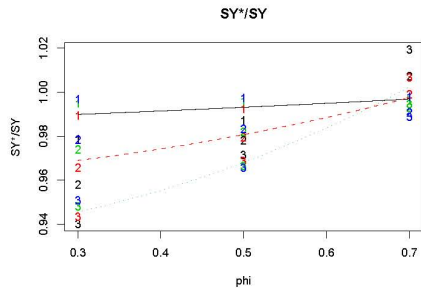
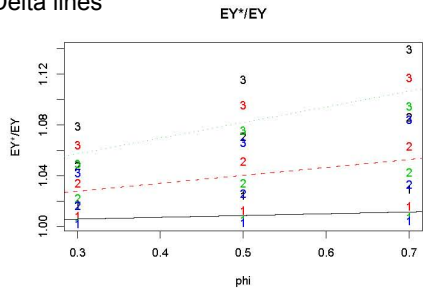
SZ^*/SZ



Moderate/Slow

Scenario C (5 periodic screens)

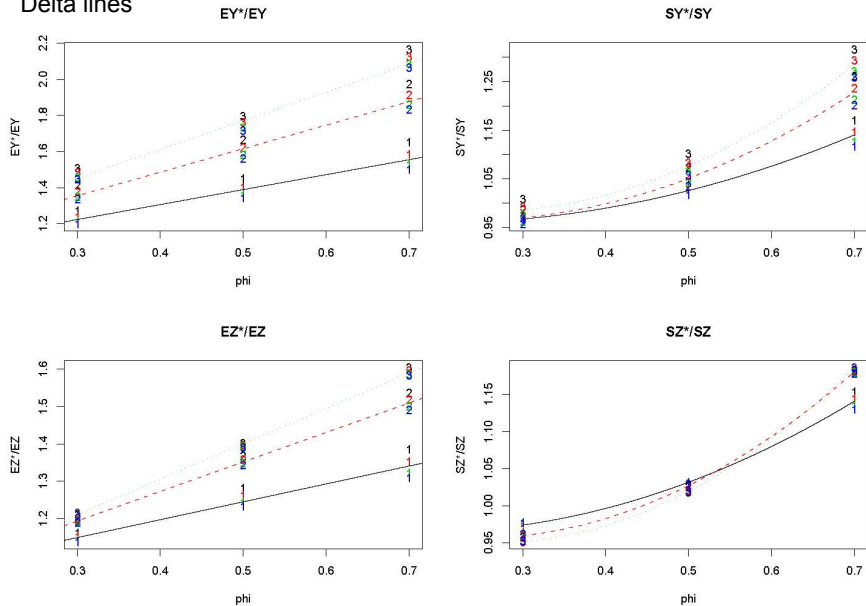
Delta lines



V. Fast/Moderate

Scenario E (5 periodic screens)

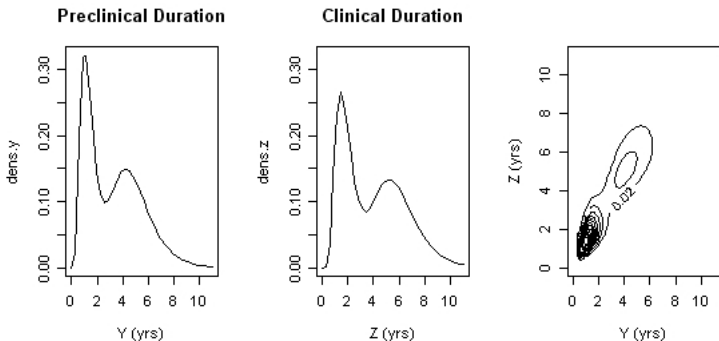
Delta lines



Variable Test Sensitivity

- Start with mixture of correlated lognormals:

$$f(y, z) = \phi \mathcal{LN}(y, z; \mu_Y = 1.5, \sigma_Y = 1, \mu_Z = 2, \sigma_Z = 1) \\ + (1 - \phi) \mathcal{LN}(y, z; \mu_Y = 5, \sigma_Y = 2, \mu_Z = 6, \sigma_Z = 2)$$



Example

- $\mathcal{LN}(y, z; \mu_Y = 1.5, \sigma_Y = 1, \mu_Z = 2, \sigma_Z = 1)$
- $\mathcal{LN}(y, z; \mu_Y = 5, \sigma_Y = 2, \mu_Z = 6, \sigma_Z = 2)$
- $\delta = 1.0, k = 4$
- As ρ increases, so does $E(Z^*)/E(Z)$

ρ	β_ϕ	$\beta_{(1-\phi)}$	$\frac{E(Z^*)}{E(Z)}$
0.70	.9	.9	1.10
0.70	.7	.7	1.11
0.70	.7	.9	1.12
0.80	.9	.9	1.16
0.80	.7	.7	1.19
0.80	.7	.9	1.21 ←

- Higher β decreases bias unless it's disproportionately higher for slow disease

5. Conclusion

- Incorporate variable test sensitivity $\beta(y)$
- Quantify size-biased sampling under various conditions and distributions
- Determine most influential factors
 - distribution
 - ratio $\delta/E(Y)$
 - test sensitivity
 - correlation between Y and Z
 - ratio $E(Y)_\phi/E(Y)_{(1-\phi)}$
- Adjust survival time for length-bias
→ unbiased estimates of screening benefit