

A Sequential Stochastic Assignment Problem with Random Number of Jobs

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joint work with

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Outline

- ▶ Sequential Stochastic Assignment Problem with Fixed Number of Jobs
- ▶ Sequential Stochastic Assignment Problem with Random Number of Jobs
- ▶ Selecting one of the k Best Values with Random Number of Alternatives

Part 1: Sequential Stochastic Assignment Problem with Fixed Number of Jobs

Classical Formulation ^a

- ▶ Suppose that n jobs arrive sequentially in time.
- ▶ The t th job, $1 \leq t \leq n$, is identified with a random variable Y_t which is observed.
- ▶ The jobs must be assigned to n persons which have known “values” p_1, \dots, p_n .
- ▶ If the t th job is assigned to the j th person then a reward of $p_j Y_t$ is obtained and the person j becomes unavailable.
- ▶ **The goal:** to maximize expected total reward

$$S_n(\pi) := E \sum_{t=1}^n p_{\pi_t} Y_t.$$

^aDerman, Lieberman & Ross (1972). A sequential stochastic assignment problem. *Management Science*, **18**, 349–355.

Notations

- ▶ Assume that Y_1, \dots, Y_n are integrable independent random variables defined on probability space (Ω, \mathcal{F}, P) .
- ▶ let F_t^a be the distribution function of Y_t , $t = 1, \dots, n$.
- ▶ Let \mathcal{Y}_t denote the σ -field generated by (Y_1, \dots, Y_t) :
 $\mathcal{Y}_t = \sigma(Y_1, \dots, Y_t)$, $1 \leq t \leq n$.
- ▶ $\pi = (\pi_1, \dots, \pi_n)$ is a permutation of $\{1, \dots, n\}$ defined on (Ω, \mathcal{F}) .
- ▶ We say that π is an assignment policy if $\{\pi_t = j\} \in \mathcal{Y}_t$ for every $1 \leq j \leq n$ and $1 \leq t \leq n$:
 π is a policy if it is non-anticipating relative to the filtration $\mathcal{Y} = \{\mathcal{Y}_t, 1 \leq t \leq n\}$ so that t th job is assigned on the basis of information in \mathcal{Y}_t .

^aAlbright, S. C., Jr. (1972). Stochastic sequential assignment problems. *Technical report*, 147, Stanford University.

Formal Statement: Problem (AP1)

- ▶ Given a vector $p = (p_1, \dots, p_n)$, with $p_1 \leq p_2 \leq \dots \leq p_n$,
- ▶ we want to maximize *the total expected reward*
 $S_n(\pi) := \mathbb{E} \sum_{t=1}^n p_{\pi_t} Y_t$ with respect to $\pi \in \Pi(\mathcal{Y})$.
- ▶ The policy π^* is called *optimal* if $S_n(\pi^*) = \sup_{\pi \in \Pi(\mathcal{Y})} S_n(\pi)$.

Useful representation:

$$\sum_{t=1}^n p_{\pi_t} Y_t = \sum_{t=1}^n \sum_{j=1}^n p_j Y_t \mathbf{1}\{\pi_t = j\} = \sum_{j=1}^n p_j Y_{\nu_j};$$

- ν_j denotes the index of the job to which the j th person is assigned: $\{\nu_j = t\} = \{\pi_t = j\}$, $1 \leq t \leq n$, $1 \leq j \leq n$.

Backward Induction Solution

► Theorem (DLR, 1972; Albright, 1972):

– There exist real numbers

$-\infty \equiv a_{0,n} \leq a_{1,n} \leq \dots \leq a_{n-1,n} \leq a_{n,n} \equiv \infty$ such that on the first step, when $Y_1 \sim F_1$ is observed, the optimal policy is

$$\pi_1^* = \sum_{j=1}^n j \mathbf{1}\{Y_1 \in (a_{j-1,n}, a_{j,n}]\}.$$

– $\{a_{j,n}\}_{j=1}^n$ do not depend on p_1, \dots, p_n and are determined by

$$a_{j,n+1} = \int_{a_{j-1,n}}^{a_{j,n}} z dF_1(z) + a_{j-1,n} F_1(a_{j-1,n}) + a_{j,n} [1 - F_1(a_{j,n})],$$

$j = 1, \dots, n$, where $-\infty \cdot 0 \equiv 0 \equiv \infty \cdot 0$.

Backward Induction Solution (con't)

- At the end of the first stage the **assigned p is removed** from the feasible set and the process repeats with the next observation, where the above calculation is then performed relative to the distribution F_2 and real numbers $-\infty \equiv a_{0,n-1} \leq a_{1,n-1} \leq \dots \leq a_{n-2,n-1} \leq a_{n-1,n-1} \equiv \infty$ are determined, and so on.

- Moreover,

$$a_{j,n+1} = EY_{\nu_j}, \quad \forall 1 \leq j \leq n,$$

i.e., $a_{j,n+1}$ is the expected value of the job which is assigned to the j th person.

Remark and Example

- ▶ By **backward induction** we determine a triangular array, where we use F_{n-t+2} to determine $\{a_{.,t}\}$:

$$a_{1,2}$$

$$a_{1,3}, a_{2,3}$$

$$a_{1,4}, a_{2,4}, a_{3,4}$$

⋮

$$a_{1,n}, a_{2,n}, \dots, a_{n-1,n}$$

$$a_{1,n+1}, a_{2,n+1}, \dots, a_{n,n+1} \Rightarrow S_n(\pi^*) = p_1 \cdot a_{1,n+1} + \dots + p_n \cdot a_{n,n+1}$$

- ▶ **Example:** $X_1 \sim X_2 \sim X_3 \sim \text{Uniform}[0, 1]$

$$a_{1,2} = 1/2$$

$$a_{1,3} = 3/8, a_{2,3} = 5/8$$

$$a_{1,4} = 39/128, a_{2,4} = 39/128, a_{3,4} = 89/128 \Rightarrow$$

$$S_3 = p_1 \cdot 39/128 + p_2 \cdot 1/2 + p_3 \cdot 89/128.$$

Part 2: Sequential Stochastic Assignment Problem with Random Number of Jobs

Problem (AP2)

- ▶ Let N be a positive integer-valued random variable with **known distribution** $\gamma = \{\gamma_k\}$, $\gamma_k = \mathbb{P}(N = k)$, $k = 1, \dots, N_{\max}$, where N_{\max} can be infinite.
- ▶ Let Y_1, Y_2, \dots be an infinite sequence of integrable independent random variables with distributions F_1, F_2, \dots , independent of N .
- ▶ Given real numbers $p_1 \leq \dots \leq p_{N_{\max}}$ the objective is to maximize the expected total reward

$$S_\gamma(\pi) = \mathbb{E} \sum_{t=1}^N p_{\pi_t} Y_t$$

over all policies $\pi \in \Pi(\mathcal{Y})$.

Random Number of Jobs

► Theorem: ^a

► In Problem (AP2) assume that $N_{\max} < \infty$ and let

$$\tilde{Y}_t := Y_t \sum_{k=t}^{N_{\max}} \gamma_k, \quad t = 1, \dots, N_{\max}$$

.

► For any $\pi \in \Pi(\mathcal{Y})$ one has

$$S_\gamma(\pi) = \mathbb{E} \sum_{t=1}^{N_{\max}} p_{\pi_t} \tilde{Y}_t,$$

and the optimal policy in Problem (AP2) coincides with the optimal policy in Problem (AP1) associated with fixed horizon $n = N_{\max}$ and job sizes $\tilde{Y}_1, \dots, \tilde{Y}_{N_{\max}}$.

^aGoldenshluger, A., Malinovsky, Y., Zeevi, A. (2019). A Unified Approach for Solving Sequential Selection Problems. *arXiv:1901.04183*.

Proof

- ▶ For any $\pi \in \Pi(\mathcal{Y})$ we have

$$S_\gamma(\pi) = \mathbb{E} \sum_{t=1}^N p_{\pi_t} Y_t = \sum_{t=1}^{N_{\max}} \mathbb{E}[p_{\pi_t} Y_t \mathbf{1}(N \geq t)],$$

- ▶ and

$$\begin{aligned} \mathbb{E}[p_{\pi_t} Y_t \mathbf{1}(N \geq t)] &= \mathbb{E} \sum_{k=t}^{N_{\max}} \mathbb{E}\left\{ [p_{\pi_t} Y_t \mathbf{1}(N = k)] \mid \mathcal{Y}_t \right\} = \mathbb{E}\left\{ p_{\pi_t} Y_t \sum_{k=t}^{N_{\max}} \gamma_k \right\} \\ &= \mathbb{E}\left\{ p_{\pi_t} \tilde{Y}_t \right\}, \end{aligned}$$

where we have used the fact that π_t is \mathcal{Y}_t -measurable, and N is independent of \mathcal{Y}_t .

- ▶ Therefore $\mathbb{E} \sum_{t=1}^N p_{\pi_t} Y_t = \mathbb{E} \sum_{t=1}^{N_{\max}} p_{\pi_t} \tilde{Y}_t$.
- ▶ Note that \tilde{Y}_t are independent random variables, and σ -fields $\tilde{\mathcal{Y}}_t$ and \mathcal{Y}_t are identical. This implies the stated result.

**Part 3: Selecting one of the k Best
Values with Random Number of
Alternatives**

Sequential Selection Problems

- ▶ Let X_1, X_2, \dots be an infinite sequence of independent identically distributed continuous random variables defined on a probability space (Ω, \mathcal{F}, P) .

▶

$$R_t := \sum_{j=1}^t \mathbf{1}(X_t \leq X_j), \quad A_{t,n} := \sum_{j=1}^n \mathbf{1}(X_t \leq X_j), \quad t = 1, \dots, n.$$

- ▶ Let $\mathcal{R}_t := \sigma(R_1, \dots, R_t)$ and $\mathcal{X}_t := \sigma(X_1, \dots, X_t)$ denote the σ -fields generated by R_1, \dots, R_t and X_1, \dots, X_t
- ▶ $\mathcal{R} = (\mathcal{R}_t, 1 \leq t \leq n)$ and $\mathcal{X} = (\mathcal{X}_t, 1 \leq t \leq n)$ are the corresponding filtrations.
- ▶ The class of all stopping times of a filtration $\mathcal{Y} = (\mathcal{Y}_t, 1 \leq t \leq n)$ will be denoted $\mathcal{T}(\mathcal{Y})$; i.e., $\tau \in \mathcal{T}(\mathcal{Y})$ if $\{\tau = t\} \in \mathcal{Y}_t$ for all $1 \leq t \leq n$.

Average Reward

- **Fixed n: Problem (A1):** Let n be a fixed positive integer, and let $q : \{1, 2, \dots, n\} \rightarrow \mathbb{R}$ be a reward function. The average reward of a stopping rule $\tau \in \mathcal{T}(\mathcal{R})$ is $V_n(q; \tau) := \mathbb{E}q(A_{\tau, n})$, and we want to find the rule $\tau_* \in \mathcal{T}(\mathcal{R})$ such that

$$V_n^*(q) := \max_{\tau \in \mathcal{T}(\mathcal{R})} V_n(q; \tau) = \mathbb{E}q(A_{\tau_*, n}).$$

- **Random N: Problem (A2):** $\gamma_k = \mathbb{P}(N = k)$, $k = 1, 2, \dots, N_{\max}$, $N \perp \{X_t, t \geq 1\}$. Let $q : \{1, 2, \dots, N_{\max}\} \rightarrow \mathbb{R}$.

$$V_\gamma(q; \tau) := \mathbb{E}[q(A_{\tau, N})\mathbf{1}(\tau \leq N)].$$

We want to find the stopping rule $\tau_* \in \mathcal{T}(\mathcal{R})$ such that

$$V_\gamma^*(q) := \max_{\tau \in \mathcal{T}(\mathcal{R})} V_\gamma(q; \tau) = V_\gamma(q; \tau_*).$$

Fixed n : Gusein-Zade Stopping Problem ^a

- ▶ **Selecting One of the k Best Values:** $q(a) = q_{\text{gz}}^{(k)}(a) := \mathbf{1}\{a \leq k\}$, and the problem is to maximize $P\{A_{\tau, n} \leq k\}$ with respect to $\tau \in \mathcal{T}(\mathcal{R})$.

- ▶ **The optimal policy:** is determined by k natural numbers

$$1 \leq \pi_1 \leq \pi_2 \leq \dots \leq \pi_k$$

and proceeds as follows: pass the first $\pi_1 - 1$ observations and among the subsequent $\pi_1, \pi_1 + 1, \dots, \pi_2 - 1$ choose the first best observation; if it does not exist then among the set of observations $\pi_2, \pi_2 + 1, \dots, \pi_3 - 1$ choose one of the two best, etc.

- ▶ **Example ($n=30, k=3$):** $\pi_1 = 11, \pi_2 = 18, \pi_3 = 24$ and

^aGusein-Zade, S. M. (1966). The problem of choice and the optimal stopping rule for a sequence of independent trials. *Theory Probab. Appl.*, **11**, 472–476.

$$\max_{\tau \in \mathcal{I}(\mathcal{R})} \mathbb{P}\{A_{\tau,30} \leq 3\} = 0.73492.$$

An Auxiliary Optimal Stopping Problem: Problem (B)

- ▶ Let Y_1, \dots, Y_n be a sequence of integrable independent real-valued random variables with corresponding distributions F_1, \dots, F_n .
- ▶ For a stopping rule $\tau \in \mathcal{T}(\mathcal{Y})$ define $W_n(\tau) := \mathbb{E}Y_\tau$. The objective is to find the stopping rule $\tau_* \in \mathcal{T}(\mathcal{Y})$ such that

$$W_n^* := \max_{\tau \in \mathcal{T}(\mathcal{Y})} \mathbb{E}Y_\tau = W_n(\tau_*) = \mathbb{E}Y_{\tau_*}.$$

DLR (1972) Solution of Problem (B)

► Consider Problem (AP1) with $p_1 = 0, p_2 = 0, \dots, p_n = 1$ and by Theorem (DLR, 1972), at step t the optimal policy assign p_n to the job Y_t only if $Y_t > a_{n-t, n-(t-1)}$ and \dots

► Let $\{b_t, t \geq 1\}$ be the sequence of real numbers defined recursively by

$$* b_1 = -\infty, b_2 = \mathbf{E}Y_n,$$

$$* b_{t+1} = \int_{b_t}^{\infty} z dF_{n-t+1}(z) + b_t F_{n-t+1}(b_t), \quad t = 2, \dots, n.$$

► Let

$$\tau_* = \min\{1 \leq t \leq n : Y_t > b_{n-t+1}\};$$

then

$$W_n^* = \mathbf{E}Y_{\tau_*} = \max_{\tau \in \mathcal{T}(\mathcal{Y})} \mathbf{E}Y_{\tau} = b_{n+1}.$$

Reduction: Problems (A1) \Rightarrow Problem (B)

- Fixed Horizon n

Let

$$I_{t,n}(r) := \sum_{a=r}^{n-t+r} q(a) \frac{\binom{a-1}{r-1} \binom{n-a}{t-r}}{\binom{n}{t}} = \mathbb{E}\{q(A_{t,n}) \mid R_t = r\}, \quad r = 1, \dots, t. \quad (1)$$

$$Y_t := I_{t,n}(R_t), \quad t = 1, \dots, n. \quad (2)$$

- **Theorem**: the optimal stopping rule τ_* solving Problem (B) with random variables $\{Y_t\}$ given in (1)–(2) also solves Problem (A1):

$$V_n(q; \tau_*) = \max_{\tau \in \mathcal{T}(\mathcal{R})} \mathbb{E}q(A_{\tau,n}) = \max_{\tau \in \mathcal{T}(\mathcal{Y})} \mathbb{E}Y_{\tau}.$$

Proof

* First we note that for any stopping rule $\tau \in \mathcal{T}(\mathcal{R})$ one has $\mathbb{E}q(A_{\tau,n}) = \mathbb{E}Y_{\tau}$, where $Y_t := \mathbb{E}[q(A_{t,n})|\mathcal{R}_t]$.

*

$$\begin{aligned}\mathbb{E}q(A_{\tau,n}) &= \sum_{k=1}^n \mathbb{E}q(A_{\tau,n})\mathbf{1}\{\tau = k\} = \sum_{k=1}^n \mathbb{E}q(A_{k,n})\mathbf{1}\{\tau = k\} \\ &= \sum_{k=1}^n \mathbb{E}\left[\mathbf{1}\{\tau = k\}\mathbb{E}\{q(A_{k,n})|\mathcal{R}_k\}\right] = \sum_{k=1}^n \mathbb{E}[\mathbf{1}\{\tau = k\}Y_k] = \mathbb{E}Y_{\tau},\end{aligned}$$

where we have used the fact that $\{\tau = k\} \in \mathcal{R}_k$. This implies that $\max_{\tau \in \mathcal{T}(\mathcal{R})} \mathbb{E}q(A_{\tau,n}) = \max_{\tau \in \mathcal{T}(\mathcal{R})} \mathbb{E}Y_{\tau}$.

* To prove the theorem it suffices to show only that

$$\max_{\tau \in \mathcal{T}(\mathcal{R})} \mathbb{E}Y_{\tau} = \max_{\tau \in \mathcal{T}(\mathcal{Y})} \mathbb{E}Y_{\tau}. \quad (3)$$

Proof (Con't)

* Clearly,

$$\mathcal{Y}_t \subset \mathcal{R}_t, \quad \forall 1 \leq t \leq n. \quad (4)$$

* Because R_1, \dots, R_n are independent random variables, and $Y_t = I_{t,n}(R_t)$, $\forall t$ we have that for any $s, t \in \{1, \dots, n\}$ with $s < t$

$$P\{G_t | \mathcal{Y}_s\} = P\{G_t | \mathcal{R}_s\}, \quad \forall G_t \in \mathcal{Y}_t. \quad (5)$$

* The statement (3) follows from (4), (5) and Theorem 5.3. ^a
This concludes the proof.

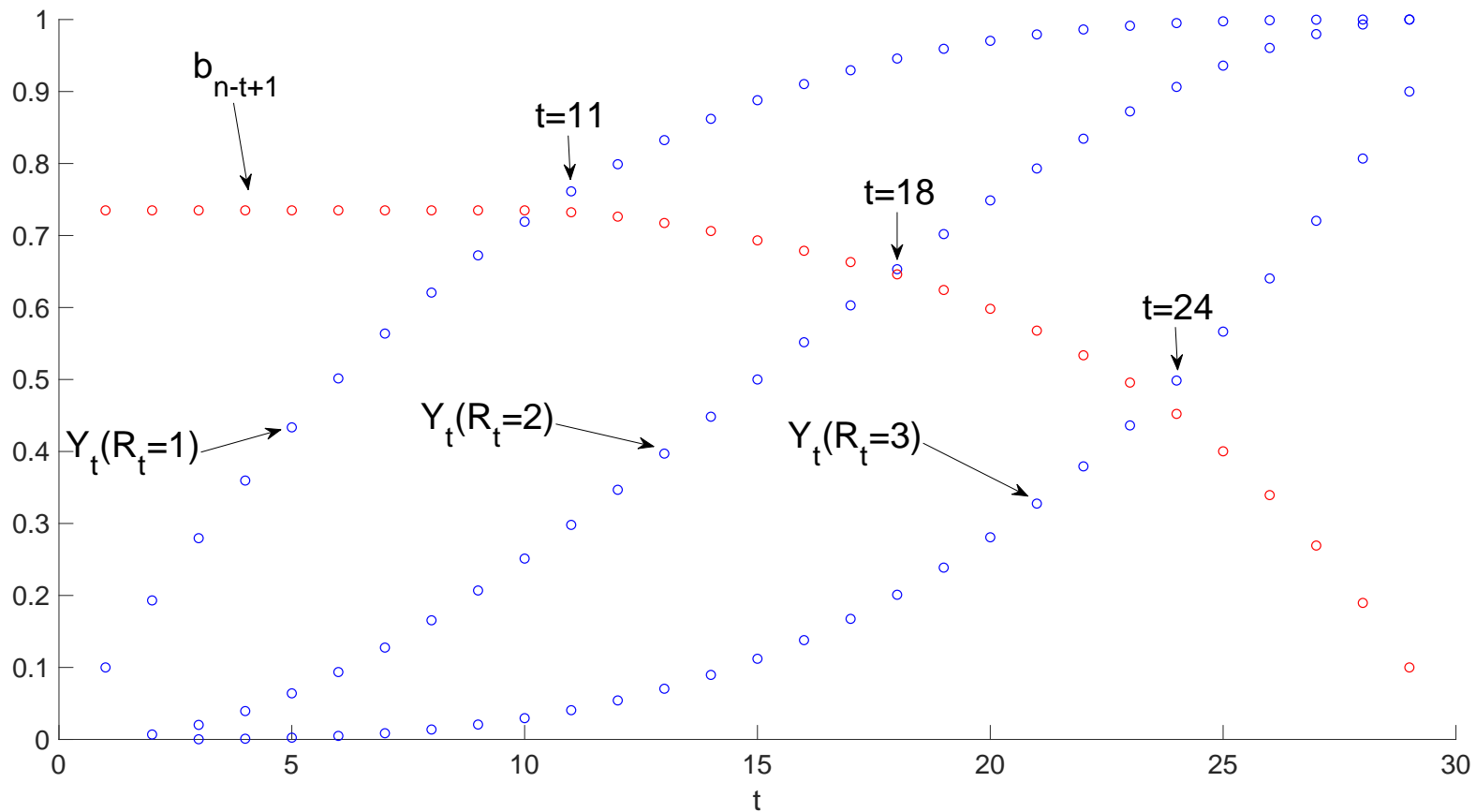
^aChow, Y. S., Robbins, H. and Siegmung, D. (1971). *Great Expectations: The Theory of Optimal Stopping*. Houghton Mifflin Company, Boston.

Numerical Values

n	k	$P(n, k)$	$E(n, k)/n$	n	k	$P(n, k)$	$E(n, k)/n$
100	2	0.57956	0.68645	1,000	2	0.57417	0.68966
	5	0.86917	0.60871		5	0.86123	0.60988
	10	0.98140	0.54236		10	0.97703	0.54434
	15	0.99755	0.50428		15	0.99609	0.50893
10,000	2	0.57363	0.68927	50,000	2	0.57358	0.68923
	5	0.86043	0.61014		5	0.86036	0.61018
	10	0.97658	0.54496		10	0.97654	0.54500
	15	0.99592	0.50947		15	0.99591	0.50950

Table 1: Optimal probabilities $P(n, k)$ and the normalized expected time elapsed until stopping $E(n, k)/n$ for selecting one of the k best values.

Optimal Strategy for $n = 30, k = 3$



$$V_n^*(q) = 0.73492$$

Reduction: Problems (A2) \Rightarrow Problem (B)

► Random Horizon

Let

$$J_t(r) := \sum_{k=t}^{N_{\max}} \gamma_k I_{t,k}(r), \quad r = 1, \dots, t. \quad (6)$$

Define

$$Y_t := J_t(R_t) = \sum_{k=t}^{N_{\max}} \gamma_k I_{t,k}(R_t), \quad t = 1, \dots, N_{\max}. \quad (7)$$

- **Theorem:** let $N_{\max} < \infty$; then the optimal stopping rule τ_* solving Problem (B) with fixed horizon N_{\max} and random variables $\{Y_t\}$ given in (6)–(7) provides the optimal solution to Problem (A2):

$$V_\gamma^*(q) = \max_{\tau \in \mathcal{I}(\mathcal{R})} V_\gamma(q; \tau) = \max_{\tau \in \mathcal{I}(\mathcal{Y})} \mathbf{E}Y_\tau = W_{N_{\max}}(\tau^*).$$

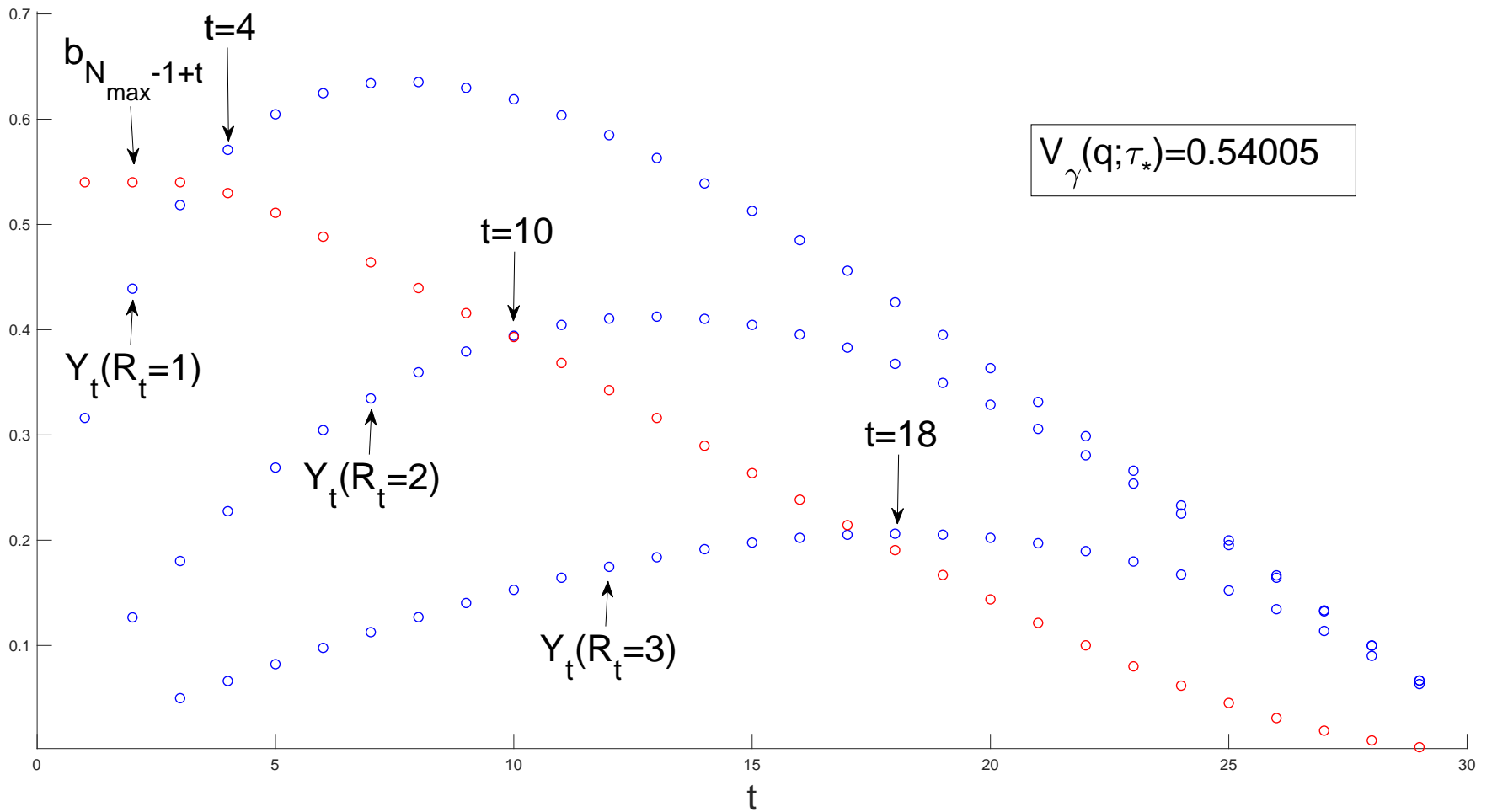
Proof

- * In Problem (A2) the reward for stopping at time t is $\tilde{q}(A_{t,N}) = q(A_{t,N})\mathbf{1}\{N \geq t\}$.

$$\begin{aligned}
 & \mathbb{E}\{q(A_{t,N})\mathbf{1}\{N \geq t\} \mid R_1 = r_1, \dots, R_{t-1} = r_{t-1}, R_t = r\} \\
 &= \sum_{k=t}^{N_{\max}} \mathbb{E}\{q(A_{t,N})\mathbf{1}\{N = k\} \mid R_1 = r_1, \dots, R_{t-1} = r_{t-1}, R_t = r\} \\
 &= \sum_{k=t}^{N_{\max}} \mathbb{E}\{\mathbf{1}\{N = k\} \mathbb{E}[q(A_{t,k}) \mid N = k, R_t = r]\} \\
 &= \sum_{k=t}^{N_{\max}} \gamma_k \sum_{a=r}^{k-t+r} q(a) \frac{\binom{a-1}{r-1} \binom{k-a}{t-r}}{\binom{k}{t}} = \sum_{k=t}^{N_{\max}} \gamma_k I_{t,k}(r) =: J_t(r). \quad (8)
 \end{aligned}$$

- * Together with (7) this implies that $\mathbb{E}\tilde{q}(A_{\tau,N}) = \mathbb{E}J_{\tau}(R_{\tau}) = \mathbb{E}Y_{\tau}$ for any $\tau \in \mathcal{T}(\mathcal{R})$. The remainder of the proof proceeds along the lines of the proof of Theorem for fixed horizon n .

Optimal Strategy for $N \sim \text{Uniform}\{1, 2, \dots, 30\}$, $k = 3$



Concluding Remarks

- ▶ The proposed framework is applicable to sequential selection problems that can be reduced to settings with independent observations and additive reward function. In particular:
 - selection problems with no-information, rank-dependent rewards and fixed or random horizon,
 - selection problems with full information when the random variables $\{X_t\}$ are observable, and the reward for stopping at time t is a function of the current observation X_t only,
 - multiple choice problems with random horizon and additive reward.

Concluding Remarks (con't)

- ▶ The proposed framework **is not applicable** to the following sequential selection problems:
 - for instance, settings with rank-dependent reward and full information as in Gnedin (2007)^a cannot be reduced to optimal stopping of a sequence of independent random variables
 - multiple choice problem with zero-one reward, where the problem of maximizing the probability of selecting k best alternatives; see, e.g., Rose (1982)^b where the problem of maximizing the probability of selecting k best alternatives was considered.

^aGnedin, A. V. (2007). Optimal stopping with rank-dependent loss. *J. Appl. Probab.*, **44**, 996–1011.

^bRose, J. S. (1972). A problem of optimal choice and assignment. *Oper. Res.*, **30**, 172–181.

Thank You !